

# Localizing Runtime Anomalies in Service-Oriented Systems

Qiang He, Xiaoyuan Xie, Yanchun Wang, Dayong Ye, Feifei Chen, Hai Jin, Yun Yang\*

**Abstract**—In a distributed, dynamic and volatile operating environment, runtime anomalies occurring in service-oriented systems (SOSs) must be located and fixed in a timely manner in order to guarantee successful delivery of outcomes in response to user requests. Monitoring all component services constantly and inspecting the entire SOS upon a runtime anomaly are impractical due to excessive resource and time consumption required, especially in large-scale scenarios. We present a spectrum-based approach that goes through a five-phase process to quickly localize runtime anomalies occurring in SOSs based on end-to-end system delays. Upon runtime anomalies, our approach calculates the similarity coefficient for each basic component (BC) of the SOS to evaluate their suspiciousness of being faulty. Our approach also calculates the delay coefficients to evaluate each BC's contribution to the severity of the end-to-end system delays. Finally, the BCs are ranked by their similarity coefficient scores and delay coefficient scores to determine the order of them being inspected. Extensive experiments are conducted to evaluate the effectiveness and efficiency of the proposed approach. The results indicate that our approach significantly outperforms random inspection and the popular Ochiai-based inspection in localizing single and multiple runtime anomalies effectively. Thus, our approach can help save time and effort for localizing runtime anomalies occurring in SOSs.

**Index Terms**—Program analysis, quality of service, Web service

## 1 INTRODUCTION

THE service-oriented computing paradigm offers an effective way to build software systems [10] that are composed of services locally or remotely accessed by an execution engine (e.g., a BPEL engine [34]). In such a service-oriented system (SOS), the component services jointly offer the functionality of the SOS and collectively fulfill its users' quality requirements.

Built from loosely coupled component services offered by independent (and often distributed) providers, SOSs operate in environments where key characteristics of the component services, such as the Quality of Service (QoS) properties, tend to be volatile. At runtime, various anomalies may occur in an SOS, e.g., unexpected workload changes, errors in the component services and failures of data transmissions, and impact on the quality of the SOS, causing end-to-end system delays. In this context, how to manage the quality of an SOS by detecting and adapting to runtime anomalies has become an important research direction [7, 10].

Response time, among various QoS dimensions, is of

particular significance in quality management for SOSs. Amazon found that every extra 100ms of latency cost them 1% in sales [43] and Google found an extra 500ms in search page generation time dropped traffic by 20% [28]. The increase in the number of time-constrained applications in the cloud, e.g., interactive and multimedia SOSs, is also driving the needs for response time management for SOSs [26]. Furthermore, the management of response time is the basis for the management of other QoS dimensions. On one hand, effective response time management promises better management of other QoS dimensions because many applications exhibit trade-offs between their response times and other QoS dimensions [30]. A video encoding application, for example, can often produce higher quality video if it is given more time to encode the video frames. On the other hand, the management of other QoS dimensions is tightly coupled with response time management. During execution, an SOS may need to be adapted to fix runtime anomalies. The adaptation itself takes time, and as a result, contributes to delaying the execution of the SOS. Thus, timely detection of runtime anomalies is significant to effective quality management for SOSs.

An intuitive solution for timely detection of runtime anomalies is to constantly monitor all the *basic components* (BCs) of an SOS, including its component services and the data transmission links (or transmissions in short) between the component services. In response to a detected anomaly, adaptation actions can be taken before performance degradation becomes noticeable by the users. However, monitoring itself may incur excessive costs [7], making it impractical to constantly monitor the entire SOS, especially in large-scale scenarios where the number of BCs of the SOS and the number of SOSs are big. To address this issue, we proposed CriMon in [16] for formu-

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lating cost-effective monitoring strategies that focus on the BCs on the critical path. However, what if an anomaly occurs in an unmonitored BC? When that happens, the unmonitored BCs must be inspected to pinpoint the anomaly. Unfortunately, a comprehensive inspection of all unmonitored BCs can be expensive and sometimes impractical due to two potential costs, i.e., *resource cost* and *system cost*. First, invocations of services may be charged if those services are owned and hosted by different organisations [50]. Even if the invocations are free, the negative impact on the SOS caused by the inspection of its BCs might further degrade its quality [7]. For example, monitoring sometimes involves sniffing network traffic and retrieving the logs of services' and SOSs' behaviours. As demonstrated in [46], those operations can result in up to 70% performance overhead, slowing down the entire SOS as a result. We identified that a quality inspection can cause as much as 40% performance overhead on a Web service under certain circumstances [21]. These issues should not exclude inspection of BCs as a promising way of detecting anomalies. Unmonitored BCs still need to be inspected to localize runtime anomalies. To achieve timely anomaly detection at minimal resource cost and system cost, the BCs that are more likely to be faulty must be prioritised in the inspection. This way, the anomalies occurring in an SOS can be quickly pinpointed so that they can be fixed in time to minimise the system delays perceived by users.

To tackle this challenge, we propose a novel approach that employs spectrum-based fault localization (SFL) technique [2] to localize runtime anomalies in SOSs. There are other fault localization techniques, e.g., program slicing [49], delta debugging [31] and model-based diagnosis [29]. The reason for our choice of SFL is that it is the most light-weight fault localization technique [40]. SFL instruments and executes a program, before it is released, with a test suite to gather program spectrum. Each program spectrum is an abstraction of an execution trace. For each test case in the test suite, the program spectrum records which program components (e.g., statements and branches) were executed and whether the test case passed or failed. Then, SFL builds a matrix that consists of two parts: the program spectrum and the binary error vector. By evaluating the *similarity* between the program spectrum and the binary error vector, a heuristic measure for each program component, which expresses the *suspiciousness* of the program component being buggy and responsible for the test-case failures. The program components can then be ordered in decreasing order of suspiciousness to provide a ranking of program components from most to least suspicious, which are provided to program developers as guidance in debugging.

Localizing runtime anomalies in an SOS is very different from localizing faults in a program written in traditional languages, such as C, Java or PHP. Anomalies in an SOS occur at runtime and must be fixed as soon as possible. Running a comprehensive test suite to localize the anomalies often results in massive calls to the component services of the SOS. It is time consuming, resource consuming, and thus impractical. Although the component services of an SOS are usually distributed, the calls to those component services are carried out and managed

by a centralised execution engine, e.g., a BPEL engine. Thus, end-to-end system delays caused by runtime anomalies can be detected by the execution engine with marginal system overhead. In response to different user requests, different parts of the SOS are activated. Thus, an SOS can be represented by multiple execution scenarios [16].

Following the methodology of SFL, upon detected end-to-end system delays, our approach generates *system spectra* that record which BCs are covered by which execution scenarios, and a binary *delay vector* that indicates which execution scenarios are experiencing delays. Then, it diagnoses the differences in the system spectra for *normal* and *delayed* execution scenarios, where normal execution scenarios are the execution scenarios with no delays and delayed execution scenarios are the execution scenarios where delays are occurring. By evaluating the similarity between the system spectrum and the binary delay vector, a *similarity coefficient* score is calculated for each BC of the SOS as a heuristic measure that expresses the *suspiciousness* of the BC being responsible for the system delay. SFL has been acknowledged as an effective approach for localizing single program faults, but it performs less favourably when localizing multiple program faults [24]. This inherent limitation of SFL automatically transmits to its application in the context of SOSs. To handle concurrent runtime anomalies, we further analyse the differentiated end-to-end delays in different execution scenarios to calculate a *delay coefficient* score for each BC to express the BC's contribution to the severity of end-to-end system delays. Next, the BCs are sorted first in decreasing order of their similarity coefficient scores and then their delay coefficient scores. Finally, the results are used to pinpoint the anomaly or, if an exact pinpoint cannot be found, presented as guidance that can reduce the SOS administrator's effort in searching for the anomaly.

The contributions of this paper are as follows:

- We propose a spectrum-based approach for localizing runtime anomalies upon end-to-end system delays. This approach does not require constant monitoring or comprehensive inspection of all BCs of the SOS.
- We propose the calculation of similarity coefficient and delay coefficient, based on which BCs can be ranked to indicate their suspiciousness of being faulty. The ranking provides system administrators with guidelines for localizing runtime anomalies.
- Extensive and comprehensive experiments are conducted to evaluate our approach using a published real-world Web service dataset, which contains over 2500 real-world Web services. The evaluation shows that our approach significantly outperforms random inspection and traditional SFL. Using our approach, the time and efforts for localizing runtime anomalies occurring in SOSs can be significantly saved.

The rest of this paper is organised as follows: Section 2 analyses the requirements with a motivating example. Section 3 describes our anomaly localization approach. Section 4 presents the experimental results to demonstrate effectiveness and efficiency of our approach. Section 5 reviews related work. Section 6 concludes the paper

and points out future work.

## 2 MOTIVATING EXAMPLE

This section presents an example SOS, namely OnlineLive, to motivate this research. As depicted in Fig. 1, this SOS offers an on-demand service to convert, subtitle and transmit live video streams. OnlineLive consists of 26 BCs.  $N_1, N_2, \dots, N_8$  represent the component services and  $E_A, E_B, \dots, E_R$  represent the data transmissions between the component services. In response to a user request, the execution process of OnlineLive is as follows:

**Step 1:**  $N_1$  splits the live media stream selected by the user into separate video and audio streams.

**Step 2:** The video and audio streams are processed in parallel. Specifically,

- For normal users,  $N_2$  encodes 360p video stream and  $N_3$  embeds advertisements into the video stream. For ad-free premium users,  $N_4$  encodes 1080p video stream.
- $N_5$  generates the subtitle by performing speech recognition on the audio stream. Then, based on the user's preference or country/region, the subtitle is sent to either  $N_6$  or  $N_7$  to be translated into one of the two optional languages.

**Step 3:**  $N_8$  produces a media stream by merging and synchronising the video stream, audio stream and translated subtitle.

**Step 4:** The media stream is transmitted to the user.

OnlineLive must process the media stream timely and continuously. Otherwise, the user will receive a jittering media stream. When anomalies occur and cause delays to OnlineLive, the BCs must be inspected to identify the faulty ones. Random inspection is a simple approach. However, we must prioritise the BCs that are more likely to be faulty in order to localize the anomalies rapidly. By doing so, adaptation actions can be taken to fix the anomaly in time to avoid or reduce the system delay caused by the anomalies.

The most intuitive solution for timely detection of runtime anomalies is to constantly monitor all the BCs of OnlineLive. Another solution is to comprehensively inspect the status and quality of all BCs every time a system delay occurs. As discussed in Section 1, both solutions are expensive in terms of time and resource consumption, and thus are often impractical. Another approach is to employ end-to-end quality information of OnlineLive for

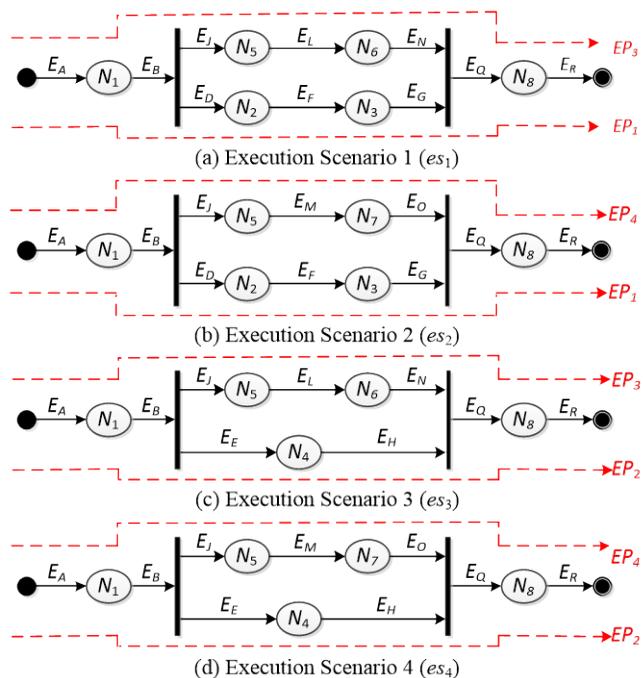


Fig. 2. Execution scenarios of OnlineLive.

localizing the occurring anomalies. Four execution scenarios can be identified from OnlineLive:  $es_1=\{EP_1, EP_3\}$ ,  $es_2=\{EP_1, EP_4\}$ ,  $es_3=\{EP_2, EP_3\}$  and  $es_4=\{EP_2, EP_4\}$ , as presented in Fig. 2. Which one will be executed in response to a user request is dependent on the runtime decisions made at the two branch structures. Assume that the constraint for the response time of OnlineLive is 3.0s but now it is taking OnlineLive 3.5s to process user requests, resulting in a system delay. System logs<sub>1</sub> indicate that the current end-to-end response times of  $es_1, es_2, es_3$  and  $es_4$  are 3.3s, 2.4s, 3.5s and 2.5s respectively. Obviously,  $es_1$  and  $es_3$  have violated the constraint of response time, thus there must be at least one faulty BC in both of these two execution scenarios. By comparing these four execution scenarios, we can find that:  $E_L, N_6$  and  $E_N$  are exclusively invoked by these two delayed execution scenarios;  $E_M, N_7$  and  $E_O$  are only involved in the normal execution scenarios; while the rest BCs are included in both cases. Therefore, we can reasonably infer that  $E_L, N_6$  and  $E_N$  are the most suspicious BCs that result in the system delay. Furthermore, multiple concurrent anomalies occurring in OnlineLive will cause a more severe end-to-end system delay, which is determined by the execution scenario with the maximum execution time among  $es_1, es_2, es_3$  and  $es_4$ . If one execution scenario is experiencing an especially severe delay, the BCs that belong to that execution scenario, especially those that exclusively belong to that execution scenario, are more likely to be faulty. Those BCs must be prioritised in the inspection.

By analysing end-to-end quality information, our approach can help the system administrator localize the faulty BCs quickly by ranking the BCs of Onlive in descending order of their suspiciousness of being faulty.

## 3 ANOMALY LOCALIZATION

Our anomaly localization approach is designed as a five-phase process as shown in Fig. 3. In Phase 1, the loops in

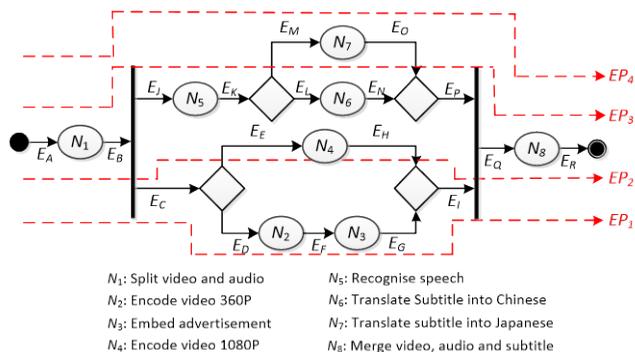


Fig. 1. Process of OnlineLive.

the SOS are peeled. In Phase 2, a set of execution scenarios are identified from the SOS. Then, in Phase 3, the system spectra are constructed. In Phase 4, similarity coefficients and delay coefficients are calculated for the BCs of the SOS. Finally, in Phase 5, the BCs are ranked based on their similarity coefficients and delay coefficients to suggest an order in which the BCs should be inspected. Details of these phases are presented in Section 3.1 to Section 3.5 respectively.

### 3.1 Phase 1: Loop Peeling

In this research, we use four types of basic compositional structures, i.e., *sequence*, *branch*, *loop* and *parallel* for representing and analysing SOS. These compositional structures are included in BPMN [35], and addressed by BPEL [34] - the de facto standard for specifying service-oriented business processes. They are also adopted in many other studies on SOS [5, 16, 50].

- **Sequence.** In a sequence structure, the BCs are executed one by one.
- **Branch.** In a branch structure, only one branch is selected for execution. For a set of branches  $\{b_1, \dots, b_n\}$ , the execution probability distribution  $\{p(b_1), \dots, p(b_n)\}$ ,  $(0 \leq p(b_i) \leq 1, \sum_{i=1}^n p(b_i) = 1.0)$  is specified, where  $p(b_i)$ ,  $i=1, \dots, n$ , is the probability that the  $i^{\text{th}}$  branch is selected for execution.
- **Loop.** In a loop structure, the loop is executed for  $n$  ( $0 \leq n \leq MNI$ ) times. For a loop, the probability distribution  $\{p_0, \dots, p_{MNI}\}$ ,  $(0 \leq p_i \leq 1, \sum_{i=0}^{MNI} p_i = 1.0)$  is specified, where  $p_i$ ,  $i=0, \dots, MNI$ , is the probability that the loop iterates for  $i$  times and  $MNI$  is the expected maximum number of iterations for the loop.
- **Parallel.** In a parallel structure, all the branches are executed at the same time.

The probabilities,  $p(b_i)$ ,  $p_i$  and the maximum number of iterations can be evaluated based on the past executions of the SOS or can be specified by the developer [5]. We assume that for a loop, the  $MNI$  can be determined or estimated. Otherwise, without an upper bound for the number of iterations, the execution times of the execution paths that contain the loop cannot be calculated since the loop may iterate infinitely.

We represent service compositions using UML activity diagrams, where the nodes represent component services and the edges represent data transmissions. Without losing generality, we assume that a service composition is characterised by only one entry point and one exit point,

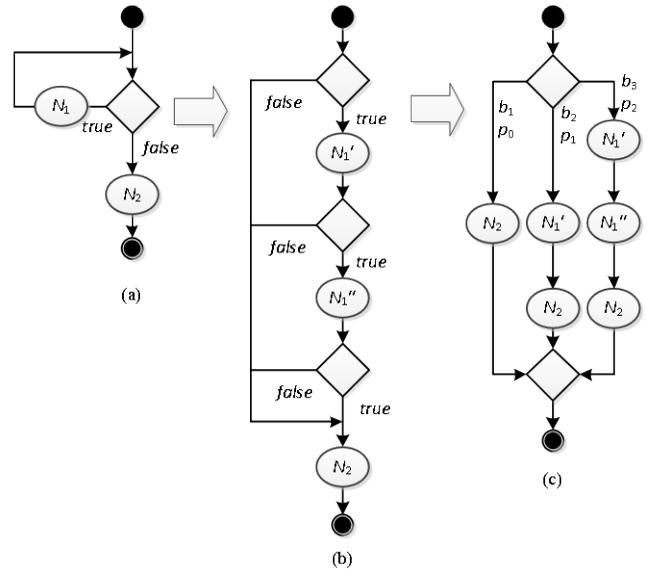


Fig. 4. Loop peeling process.

and only includes *structured loops* with only one entry point and one exit point. If a service composition includes loops, we peel the loops by representing loop iterations as a set of branches with corresponding execution probabilities [5]. Fig. 4 gives an example of peeling a loop structure ( $MNI=2$ ) by transforming it into a branch structure that contains three branches  $b_1$ ,  $b_2$  and  $b_3$ , where  $p_0$ ,  $p_1$  and  $p_2$  are the probabilities that  $b_1$ ,  $b_2$  and  $b_3$  are selected for execution respectively. (Note that the first branch  $b_1$  is selected if the loop iterates for 0 times, i.e., corresponding to  $p_0$ ).

### 3.2 Phase 2: Execution Scenario Identification

In a service composition where branches or loops are involved, different execution paths may be selected for execution. Thus, multiple *execution scenarios* can be identified from the service composition. These execution scenarios do not contain branch or loop structures. As depicted in Fig. 2, four execution scenarios can be identified from OnlineLive:  $es_1=\{EP_1, EP_3\}$ ,  $es_2=\{EP_1, EP_4\}$ ,  $es_3=\{EP_2, EP_3\}$  and  $es_4=\{EP_2, EP_4\}$ . The identified execution scenarios will be used for constructing the system spectra as described next in Section 3.3.

### 3.3 Phase 3: System Spectra Construction

Anomaly localization requires analysis of the differences in system spectra [42] for individual execution scenarios of the SOS. A system spectrum is a collection of execution traces that shows which BCs were included during an execution of the SOS. In this research, a system spectrum contains a flag for each BC of the SOS to indicate whether

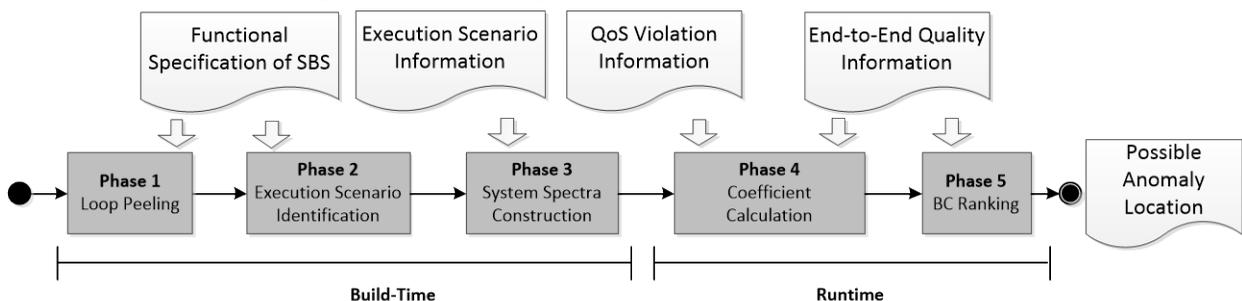


Fig. 3. Anomaly localization approach.

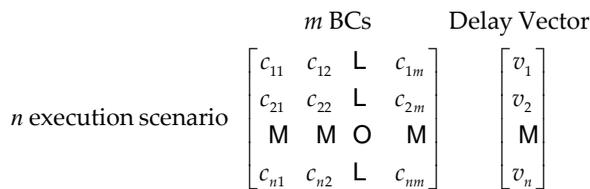


Fig. 5. System spectra and delay vector.

or not that BC is included in a particular execution scenario of the SOS. Take OnlineLive as an example. There are four system spectra, one for each execution scenario, as presented by the four dark grey columns in Table 1.

The system spectra of  $n$  execution scenarios constitute an  $n \times m$  binary matrix, where  $m$  is the number of BCs in an SOS, as presented in Fig. 5. Given  $i \in [1, \dots, n]$  and  $j \in [1, \dots, m]$ ,  $c_{ij}=1$  indicates that the  $j^{\text{th}}$  BC is included in the  $i^{\text{th}}$  execution scenario. The four dark grey columns in Table 1 constitute the system spectra of OnlineLive. Each row in the matrix is a system spectrum that indicates whether a BC is part of an execution scenario. For example, the first row, (1, 1, 1, 1), indicates that  $N_1$  is part of all four execution scenarios,  $es_1$ ,  $es_2$ ,  $es_3$  and  $es_4$ . The second row, (1, 1, 0, 0) indicates that  $N_2$  is part of  $es_1$  and  $es_2$ , but not  $es_3$  and  $es_4$ . The system spectra will be used in the similarity coefficient calculation for anomaly localization discussed next in Section 3.4.

### 3.4 Phase 4: Coefficient Calculation

This section presents the calculation of two coefficients, similarity coefficient and delay coefficient. They indicate

the suspiciousness of BCs being faulty and are used in Phase 5 for ranking the BCs for inspection.

#### 3.4.1 Similarity Coefficient

A similarity coefficient score is calculated based on a BC's involvement in all the normal and delayed execution scenarios. It indicates the suspiciousness of a BC being faulty. It also approximates the reliability of a BC [41], where a high similarity coefficient score implies low reliability.

In Fig. 5, the binary delay vector that contains  $n$  flags is used to represent which execution scenarios are experiencing a delay in processing user requests. Given  $i \in [1, \dots, n]$ ,  $v_i=1$  indicates that  $es_i$  is experiencing a delay, and 0 otherwise. To complete the delay vector, the time consumed by each execution scenario in processing user requests needs to be logged. It can be obtained from the execution engine of the SOS by calculating the time difference between the arrival of user requests and the delivery of corresponding outcomes. This is feasible because the execution engine that invokes the distributed services is centrally managed by the SOS provider [38]. Each request will traverse one execution scenario, depending on the runtime decisions made in the branch structures of the SOS. Which execution scenario each user request traverses also needs to be recorded to complete the delay vector depicted in Fig. 5. This can be achieved by recording the runtime decisions made at the branches by the execution engine of the SOS. Now assume that an anomaly is occurring in  $N_6$ , slowing it down and violating the constraint for the response time of OnlineLive. Because  $N_6$

TABLE 1  
SYSTEM SPECTRA, SIMILARITY COEFFICIENT AND BC RANKS FOR ONLINELIVE  
(ANOMALY OCCURRING IN  $N_6$ )

Basic Component	Execution Scenario				$n_{11}$	$n_{10}$	$n_{01}$	$n_{00}$	Similarity Coefficient			Ranking (by $c_o$ )
	$es_1$	$es_2$	$es_3$	$es_4$					$c_I$	$c_T$	$c_O$	
$N_1$	1	1	1	1	2	2	0	0	0.50	0.50	0.71	11
$N_2$	1	1	0	0	1	1	1	1	0.33	0.50	0.50	19
$N_3$	1	1	0	0	1	1	1	1	0.33	0.50	0.50	19
$N_4$	0	0	1	1	1	1	1	1	0.33	0.50	0.50	19
$N_5$	1	1	1	1	2	2	0	0	0.50	0.50	0.71	11
$N_6$	1	0	1	0	2	0	0	2	1.00	1.00	1.00	3
$N_7$	0	1	0	1	0	2	2	0	0.00	0.00	0.00	22
$N_8$	1	1	1	1	2	2	0	0	0.50	0.50	0.71	11
$E_A$	1	1	1	1	2	2	0	0	0.50	0.50	0.71	11
$E_B$	1	1	1	1	2	2	0	0	0.50	0.50	0.71	11
$E_D$	1	1	0	0	1	1	1	1	0.33	0.50	0.50	19
$E_E$	0	0	1	1	1	1	1	1	0.33	0.50	0.50	19
$E_F$	1	1	0	0	1	1	1	1	0.33	0.50	0.50	19
$E_G$	1	1	0	0	1	1	1	1	0.33	0.50	0.50	19
$E_H$	0	0	1	1	1	1	1	1	0.33	0.50	0.50	19
$E_J$	1	1	1	1	2	2	0	0	0.50	0.50	0.71	11
$E_L$	1	0	1	0	2	0	0	2	1.00	1.00	1.00	3
$E_M$	0	1	0	1	0	2	2	0	0.00	0.00	0.00	22
$E_N$	1	0	1	0	2	0	0	2	1.00	1.00	1.00	3
$E_O$	0	1	0	1	0	2	2	0	0.00	0.00	0.00	22
$E_Q$	1	1	1	1	2	2	0	0	0.50	0.50	0.71	11
$E_R$	1	1	1	1	2	2	0	0	0.50	0.50	0.71	11
Delay	1	0	1	0								

is included in  $EP_3$ , the system delay is only observed in  $es_1$  and  $es_3$ . The system delay vector will be:  $(1, 0, 1, 0)$ . Table 1 presents the system spectra constructed for OnlineLive.

The similarity between the binary delay vector and a system spectrum indicates the suspiciousness of the corresponding BC being responsible for the occurring system delay. In Table 1, only the system spectra corresponding to  $N_6$ ,  $E_L$  and  $E_N$  are consistent with the delay vector – they are all  $(1, 0, 1, 0)$ . Thus,  $N_6$ ,  $E_L$  and  $E_N$  can be identified as the most suspicious BCs and should be inspected with the highest priority.

For large-scale applications, the system spectra and the delay vectors can be remarkably large and complex. The suspicious BCs are those whose corresponding system spectra are most statistically similar to the delay vector. In order to statistically analyse the similarity between the system spectra and the delay vector, we now extract the features of the raw data in the system spectra and construct a *feature matrix* by comparing the system spectra with the delay vector. The feature matrix has four columns for representing four features:

$$n_{pq}(S_j) = |\{i \mid x_{ij} = p \wedge v_i = q\}|, p, q \in \{0,1\} \quad (1)$$

where  $x_{ij}=1$  indicates that BC  $S_j$  is included in  $es_i$ , and 0 otherwise,  $v_i=1$  indicates that a delay is occurring in  $es_i$  (i.e., a delayed execution scenario) and 0 (a normal execution scenario) otherwise.  $n_{10}(S_j)$  and  $n_{11}(S_j)$  are the numbers of normal and delayed execution scenarios respectively where  $S_j$  is involved.  $n_{00}(S_j)$  and  $n_{01}(S_j)$  are the numbers of normal and delayed execution scenarios respectively where  $S_j$  is NOT involved.

Now the statistical similarity between each of the system spectrum and the delay vector can be quantified by calculating the similarity coefficients of the BCs. Calculated in different ways, many similarity coefficients exist, e.g., Jaccard coefficient  $c_j$ , used by the Pinpoint tool [12], the coefficient  $c_T$  used in the Tarantula fault localization tool [25] and the Ochiai coefficient  $c_O$  often used in the molecular biology domain [14]:

$$c_j(S_j) = \frac{n_{11}(S_j)}{n_{11}(S_j) + n_{01}(S_j) + n_{10}(S_j)} \quad (2)$$

$$c_T(S_j) = \begin{cases} 1.00 & \text{if } n_{10}(S_j) + n_{00}(S_j) = 0 \\ \frac{n_{11}(S_j)}{n_{11}(S_j) + n_{01}(S_j)} & \text{otherwise} \\ \frac{n_{11}(S_j)}{n_{11}(S_j) + n_{01}(S_j)} + \frac{n_{10}(S_j)}{n_{10}(S_j) + n_{00}(S_j)} & \end{cases} \quad (3)$$

$$c_O(S_j) = \frac{n_{11}(S_j)}{\sqrt{(n_{11}(S_j) + n_{01}(S_j)) \times (n_{11}(S_j) + n_{10}(S_j))}} \quad (4)$$

where  $S_j$  is a BC. A higher similarity coefficient indicates higher suspiciousness of a BC being faulty.

### 3.4.2 Delay Coefficient

In a volatile operating environment, multiple anomalies may occur at the same time to different BCs of an SOS, especially for large-scale SOSs that consist of many BCs.

Multiple concurrent anomalies usually cause more severe delays than a single anomaly. As discussed in Section 2, the BCs in an execution scenario where a more severe delay is occurring are more suspicious of being faulty. However, it is not reflected by the Jaccard coefficient, Tarantula coefficient and Ochiai coefficient because they take into account only whether there is a delay in an execution scenario. The binary delay vector does not indicate the severity of the delay occurring in each of the execution scenarios. When multiple anomalies occur at the same time, the probability is high that most, sometimes even all, of the execution scenarios will be delayed, leading to a delay vector with many 1s, even full 1s, i.e.,  $(1, 1, \dots, 1)$ . Such a delay vector will lower the variety of the feature vectors (the light grey rows) in the feature matrix, making it harder to use the similarity coefficients to distinguish the BCs from each other.

Take Fig. 2 for example: suppose two anomalies are occurring, one to service  $N_5$  and the other to service  $N_6$ . Table 2 presents the corresponding system spectra, similarity coefficients and BC ranks. As presented by the corresponding system spectra,  $N_5$  belongs to all four execution scenarios and  $N_6$  belongs to  $es_1$  and  $es_3$ .  $N_5$  causes a delay to all four execution scenarios, resulting in a delay vector  $(1, 1, 1, 1)$ . Constructed based on the comparison between the system spectra and the all-1 delay vector, the feature matrix contains only  $(4, 0, 0, 0)$  and  $(2, 2, 0, 0)$ . The consequence is low variety in the similarity coefficients and rankings for the BCs. As presented by the *Ranking (by  $c_j$ )* column, besides  $N_5$ , seven other BCs are ranked 8 as the most suspicious BCs.  $N_6$  is ranked 21 together with all the other BCs. In the best case scenario,  $N_5$  is pinpointed as a fault BC at the first inspection of all the BCs ranked 8 and  $N_6$  is pinpointed as an abnormal BC at the second inspection. The total number of BCs that must be inspected to pinpoint  $N_5$  and  $N_6$  is 2. In the worse case scenario, the number is 21. On average, it is 11.5, getting close to random inspection, which needs to inspect an average of 16.18 BCs to pinpoint both  $N_5$  and  $N_6$ . Apparently, the traditional SFL techniques based on Jaccard coefficient, Tarantula coefficient and Ochiai coefficient cannot handle concurrent anomalies effectively.

In order to address this issue, we need to take into account the severity of the delays caused to each execution scenario. In Fig. 2, concurrent anomalies occurring in  $N_5$  and  $N_6$  will cause more severe delays to  $es_1$  and  $es_3$  compared to  $es_2$  and  $es_4$  because  $es_1$  and  $es_3$  contain both  $N_5$  and  $N_6$  while  $es_2$  and  $es_4$  contain only  $N_5$ . We calculate the *delay coefficient*, denoted by  $c_d$ , to further indicate the suspiciousness of a BC based on the severity of the delays caused by anomalies to the execution scenarios. Given a set of  $res_j(es_i)$ ,  $j=1, \dots, h$ , logged as the response time of execution scenario  $es_i$  for processing  $n$  user requests after the system delay is detected, we calculate the standard deviation of  $res_j(es_i)$  ( $j=1, \dots, h$ ) from its normal value  $res(es_i)$  :

$$SD(res(es_i)) = \sqrt{\frac{1}{h} \sum_{j=1}^h (res_j(es_i) - res(es_i))^2} \quad (5)$$

We measure the severity of a delay by measuring the

TABLE 2  
SYSTEM SPECTRA, SIMILARITY COEFFICIENT AND BC RANKS FOR ONLINELIVE  
(ANOMALIES OCCURING IN  $N_5$  AND  $N_6$ )

Basic Component	Execution Scenario				$n_{11}$	$n_{10}$	$n_{01}$	$n_{00}$	Similarity Coefficient			Ranking (by $c_o$ )	Delay Coefficient $c_D$	Ranking (by $c_D$ )
	$es_1$	$es_2$	$es_3$	$es_4$					$c_j$	$c_T$	$c_o$			
$N_1$	1	1	1	1	4	0	0	0	1	1	1	8	0.47	8
$N_2$	1	1	0	0	2	0	2	0	0.5	0.5	0.71	22	0.55	16
$N_3$	1	1	0	0	2	0	2	0	0.5	0.5	0.71	22	0.55	16
$N_4$	0	0	1	1	2	0	2	0	0.5	0.5	0.71	22	0.39	19
$N_5$	1	1	1	1	4	0	0	0	1	1	1	8	0.47	8
$N_6$	1	0	1	0	2	0	2	0	0.5	0.5	0.71	22	0.89	11
$N_7$	0	1	0	1	2	0	2	0	0.5	0.5	0.71	22	0.05	22
$N_8$	1	1	1	1	4	0	0	0	1	1	1	8	0.47	8
$E_A$	1	1	1	1	4	0	0	0	1	1	1	8	0.47	8
$E_B$	1	1	1	1	4	0	0	0	1	1	1	8	0.47	8
$E_D$	1	1	0	0	2	0	2	0	0.5	0.5	0.71	22	0.55	16
$E_E$	0	0	1	1	2	0	2	0	0.5	0.5	0.71	22	0.39	19
$E_F$	1	1	0	0	2	0	2	0	0.5	0.5	0.71	22	0.55	16
$E_G$	1	1	0	0	2	0	2	0	0.5	0.5	0.71	22	0.55	16
$E_H$	0	0	1	1	2	0	2	0	0.5	0.5	0.71	22	0.39	18
$E_J$	1	1	1	1	4	0	0	0	1	1	1	8	0.47	8
$E_L$	1	0	1	0	2	0	2	0	0.5	0.5	0.71	22	0.89	11
$E_M$	0	1	0	1	2	0	2	0	0.5	0.5	0.71	22	0.05	22
$E_N$	1	0	1	0	2	0	2	0	0.5	0.5	0.71	22	0.89	11
$E_O$	0	1	0	1	2	0	2	0	0.5	0.5	0.71	22	0.05	22
$E_Q$	1	1	1	1	4	0	0	0	1	1	1	8	0.47	8
$E_R$	1	1	1	1	4	0	0	0	1	1	1	8	0.47	8
<b>Delay</b>	1	1	1	1										

standard deviation of the response times of execution scenarios. To compare the severity of delays caused by the anomalies to different execution scenarios  $es_i$  ( $i=1, \dots, m$ ) on the same scale independently of their original response times, we calculate the coefficient of variation for each  $es_i$ :

$$c_v(es_i) = \frac{SD(res(es_i))}{res(es_i)} = \frac{\sqrt{\frac{1}{h} \sum_{j=1}^h (res_j(es_i) - \overline{res(es_i)})^2}}{res(es_i)} \quad (6)$$

Next, the min-max normalisation technique, which has been widely used in a number of approaches for QoS evaluation [5, 20, 48], is employed to normalise the coefficients of variation as follows:

$$\hat{c}_v(es_i) = \begin{cases} \frac{c_v(es_i) - c_v^{min}(es_i)}{c_v^{max}(es_i) - c_v^{min}(es_i)} & \text{if } c_v^{max}(es_i) \neq c_v^{min}(es_i) \\ 1 & \text{if } c_v^{max}(es_i) = c_v^{min}(es_i) \end{cases} \quad (7)$$

where  $c_v^{max}(es_i)$  and  $c_v^{min}(es_i)$  are the maximum and minimum values in  $c_v(es_i)$ ,  $i=1, \dots, m$ . Suppose two anomalies are occurring, one to  $N_5$  and the other to  $N_6$  in Fig. 2. Table 3 presents the calculation of the normalised coefficients of variation for  $es_1, es_2, es_3$  and  $es_4$ .

The BCs that belong to an execution scenario with a severe delay (indicated by a high  $c_v$ ) should be prioritised in the inspection. However, a BC can belong to multiple execution scenarios. Thus, we average the coefficients of variation for all the execution scenarios that a BC  $S_j$  be-

longs to to calculate the delay coefficient of the BC:

$$c_D(S_j) = \frac{\sum_{i=1}^n x_{ij} \cdot \hat{c}_v(es_i)}{\sum_{i=1}^n x_{ij}} \quad (8)$$

where  $x_{ij}$  is 1 if  $S_j$  belongs to  $es_i$  and 0 otherwise.

The Delay Coefficient  $c_D$  column in Table 2 presents the delay coefficients of the BCs of OnlineLive calculated based on Table 3.

TABLE 3  
DELAY COEFFICIENT CALCULATION  
(ANOMALIES OCCURING IN  $N_5$  AND  $N_6$ )

	$es_1$	$es_2$	$es_3$	$es_4$
$\overline{res(es)}$	2.30	2.40	2.50	2.50
$res_1(es)$	4.82	3.46	4.77	3.46
$res_2(es)$	4.85	3.39	4.82	3.41
$res_3(es)$	4.75	3.41	4.83	3.39
$res_4(es)$	4.73	3.50	4.76	3.32
$SD(res(es))$	2.49	1.04	2.30	0.90
$c_v(es)$	1.08	0.43	0.92	0.36
$\hat{c}_v(es)$	1.00	0.10	0.77	0.00

### 3.5 Phase 5: BC Ranking

Based on the similarity coefficient scores, the BCs can be ranked from high to low, where the highest indicates the most suspicious. When multiple BCs share a same coefficient score, we use the approach reported in [44]: all tied BCs get the greatest ranking number. The *Ranking (by  $c_O$ )* columns in Table 1 and Table 2 present the rankings of all BCs in OnlineLive in the cases of one anomaly (occurring in  $N_6$ ) and two anomalies (occurring in  $N_5$  and  $N_6$ ) respectively. BCs can also be ranked by their delay coefficient scores. Based on their similarity coefficients and delay coefficients, we can rank BCs in different ways. One option is to first rank them by their similarity coefficient scores and then rank the ones sharing the same similarity coefficient scores by their delay coefficient scores. Table 2 presents the BCs of OnlineLive ranked in this way. 14 BCs, are ranked 22 by their similarity coefficient scores, as indicated by column *Ranking (by  $c_O$ )*. They are further ranked 11, 16, 19 and 22 by their delay coefficient scores, as indicated by column *Ranking (by  $c_D$ )*. The ranking provides system administrator with an order in which the BCs are inspected. For OnlineLive, starting from the BC with the highest ranking, 2 BCs in the best case scenario and 11 BCs in the worst case scenario need to be inspected to pinpoint both faulty BCs, i.e.,  $N_5$  and  $N_6$ . On average, 6.5 BCs need to be inspected to pinpoint both  $N_5$  and  $N_6$ , requiring much less BCs to be inspected to localize all faulty BCs compared to random inspection which needs to inspect an average of 16.18 BCs.

The BCs can also be ranked in other ways. For example, they can be ranked first by their delay coefficient scores then by their similarity coefficient scores. The similarity coefficient and delay coefficient can also be combined into one coefficient for ranking the BCs. For example, BCs can be ranked by the average of their similarity coefficient scores and delay coefficient scores, i.e.,  $(c_O+c_D)/2$ . In Section 4, we will experimentally evaluate the three ranking methods discussed above.

## 4 EXPERIMENTS

We conducted a range of experiments in a simulated volatile environment to evaluate our anomaly localization approach. Section 4.1 describes the setup of the experiments and Section 4.2 presents the experimental results.

### 4.1 Experimental Setup

For the evaluation, we have implemented the following inspection approaches:

- *Random inspection.* In response to any anomalies, this approach inspects the BCs randomly to localize the faulty BCs. It stops when all anomalies are localized. Inspected BCs are not reinspected. This inspection approach is the baseline for the evaluation.
- *$c_O$  Ranking.* This inspection approach was proposed and evaluated by us in [18]. In response to runtime anomalies, this approach evaluates the suspiciousness of a BC by calculating the Ochiai-based similarity coefficients of the BCs. The BC with highest similarity coefficient is inspected first. This approach also stops when all anomalies are localized. As discussed in Section 3.4,

there are other coefficients that can be used to calculate the similarity of BCs, e.g., Jaccard and Tarantula. We adopt the Ochiai coefficient because it outperforms the Jaccard and Tarantula coefficients in localizing runtime anomalies in SOSs as demonstrated in [18], as well as [2, 44] for traditional software programs. This approach has also been adopted and studied by many researchers on fault localization for traditional software programs [6, 32, 37, 44].

- *$c_O$ - $c_D$  Ranking.* This approach ranks the BCs first by their Ochiai-based similarity coefficient scores then by their delay coefficient scores. The BCs are then inspected by decending order of their rankings.
- *$c_D$ - $c_O$  Ranking.* This approach is similar to the  $c_O$ - $c_D$  approach, but it ranks BCs first by their delay coefficient scores then by their Ochiai-based similarity coefficient scores.
- *$(c_D+c_O)/2$  Ranking.* This approach ranks BCs by the average of their similarity coefficient scores and delay coefficient scores.

To measure the effectiveness of the anomaly localization approaches, we use the *localization cost*, i.e., the percentage of BCs of an SOS that must be inspected before localizing all faulty BCs. This measure has been widely used by other researchers [13, 44] for fault localization studies. A *low localization cost indicates high effectiveness and vice versa*. An approach with less localization cost can pinpoint the faulty BCs quicker and incurs less inspection expense, i.e., the costs for inspecting BCs during the localization process. For example, in Table 1, the ranking of  $N_6$  (the faulty BC) is 3, the same as  $E_L$  and  $E_N$ . The system administrator can choose to inspect all the suspicious BCs at the same time or to inspect them in a random order. In a worst case scenario, the cost of pinpointing the anomaly is  $3/22$ , or 13.64%. In the case of multiple anomalies, e.g., Table 2, our approach will inspect  $N_1, N_5, N_8, E_A, E_B, E_J, E_Q$  and  $E_R$  first, which are ranked 8 by their delay coefficients  $C_D$ . After that,  $N_6, E_L$  and  $E_N$ , which are ranked 11, are inspected. Thus, in the worst case scenario, the cost of pinpointing the faulty  $N_5$  and  $N_6$  with our approach is  $(8+3)/22=11/22$ , or 50.00%.

In order to evaluate our approach on different scales, we conduct 9 sets of experiments, where the component services of the SOS increases from 20 to 100 in steps of 10. Given a specific number of component services, the SOS is randomly structured with the sequence, branch and parallel structures discussed in Section 3.1. The loop structure is omitted because it will be transformed into branch structure anyway, as discussed in Section 3.1.

The response times of the BCs are generated according to different normal distributions based on a publicly available Web service dataset QWS [3, 4]. QWS comprises measurements of 9 QoS attributes (including response time) of over 2500 real-world Web services. The information about the services was collected from public UDDI registries, search engines and service portals. Their QoS values were measured using benchmark tools. At runtime, we generated and introduced a number of concurrent anomalies to randomly picked BCs to simulate a volatile operating environment. We increased the number of concurrent anomalies from 1 to 10 in steps of 1 in each

set of experiments to simulate increasing levels of volatility in the operating environment. When anomalies occurred in BCs, delays that were randomly generated according to a normal distribution were applied to the corresponding BCs. In the real world, the delays caused by different anomalies to different BCs are usually different. There are short, medium and long ones, which delay the system to varying degrees. The long ones can significantly delay the system execution while the short ones can be neglected. There are many uncertain factors that may impact on the actual delay caused to a BC by an anomaly, e.g., the severity of the anomaly and the robustness of the BC. To simplify the application of delays and provide a generalised reference in the experiments, we generated delays caused by anomalies to different BCs according to a same normal distribution. This way, we consider or focus on only the “cost saving” effects from early anomaly localization enabled by our approach.

Upon runtime anomalies, we adopted *random inspection*, *c<sub>0</sub> ranking approach* and our approaches to pinpoint the faulty BCs. All experiments were conducted on a machine with Intel i5-4570 CPU 3.20GHz and 8 GB RAM, running Windows 7 x64 Enterprise. For each set of experiment, we averaged the results obtained from 100 runs.

## 4.2 Experimental Results

### 4.2.1 Effectiveness

Table 4 presents the average localization costs of the 5 approaches across all 9 sets of experiments. On average, our three approaches, i.e., *c<sub>0</sub>-c<sub>D</sub>*, *c<sub>D</sub>-c<sub>0</sub>* and  $(c_D+c_0)/2$ , significantly outperform random inspection, by 13.1%, 18.2% and 18.1% respectively, compared to *c<sub>0</sub>* ranking approach by 6.4%. In general, the advantages of our approaches over random inspection gradually decrease as the environmental volatility (indicated by the increasing number of concurrent anomalies) increases. However, their advantages over *c<sub>0</sub>* ranking approach increase as the experimental volatility increases.

Fig. 6 compares the localization cost of our approaches (represented by *CO-CD*, *CD-CO* and  $(CO+CD)/2$ ), with random inspection (represented by *RANDOM*) and the *c<sub>0</sub>* ranking approach (represented by *CO*). In the case of a single anomaly, as expected, random inspection has to inspect approximately 50% of the BCs to pinpoint the faulty BC. In such scenarios, a single runtime anomaly does not lead to mutually-distinguishable end-to-end delays to different execution scenarios, making it impossible for our approaches to evaluate and utilise the severity of the end-to-end delays. As expected, *c<sub>0</sub>* ranking approach and our approach demonstrate similar localization cost, between 12% and 27%, outperforming random inspection by significant margins.

As the number of concurrent anomalies increases, more BCs must be inspected to pinpoint all runtime anomalies. Accordingly, the localization costs of all approaches increase as the number of anomalies increases, rapidly at the beginning and gradually afterwards. As the number of concurrent anomalies increases, the probability that all the execution scenarios experience delays increases. Delays occurring in all execution scenarios will lead to all-1 delay vectors, making it hard for *c<sub>0</sub>* ranking

TABLE 4  
AVERAGE LOCALIZATION COSTS  
(ACROSS ALL SETS OF EXPERIMENTS)

# of Anomalies	Random	<i>c<sub>0</sub></i>	<i>c<sub>0</sub>-c<sub>D</sub></i>	<i>c<sub>D</sub>-c<sub>0</sub></i>	$\frac{c_0 + c_D}{2}$
1	0.499	0.167	0.163	0.161	0.165
2	0.672	0.525	0.491	0.492	0.490
3	0.748	0.687	0.632	0.571	0.571
4	0.806	0.771	0.699	0.626	0.627
5	0.838	0.815	0.738	0.663	0.664
6	0.864	0.847	0.759	0.688	0.691
7	0.879	0.869	0.786	0.720	0.721
8	0.893	0.886	0.801	0.740	0.742
9	0.903	0.897	0.814	0.759	0.760
10	0.913	0.908	0.826	0.775	0.776
<b>Average</b>	0.802	0.737	0.671	0.619	0.621

approach to pinpoint all the faulty BCs. Using not only the similarity coefficient but also the delay coefficient, our approaches can handle concurrent anomalies more efficiently, but still requires over 70% of the BCs to be inspected when there are 7 or more concurrent anomalies.

An interesting observation is that, as the number of concurrent anomalies increases, the localization cost of *c<sub>0</sub>* ranking approach gradually approximates that of the random inspection. The localization cost of *c<sub>0</sub>* ranking approach catches up with that of the random inspection at certain points in all 9 sets of experiments, e.g., 6 in the first set of experiments, 7 in the fifth set and 8 in the ninth set. More concurrent anomalies indicate a more volatile operating environment and consequently a less reliable SOS. Thus, this observation shows that *c<sub>0</sub>* ranking approach loses its advantage over random inspection in very volatile operating environments. However, our approaches still demonstrate significantly better effectiveness than random inspection. Compared with *c<sub>0</sub>* ranking, our approaches achieve similar localization costs in relatively stable operating environments where only one or two concurrent anomalies are occurring in the SOS. However, their advantages over *c<sub>0</sub>* ranking start to manifest as the number of concurrent anomalies increases. This is clearly demonstrated by the large gap below the red line in the cases of 2+ anomalies. Compared with *c<sub>0</sub>* ranking, our approaches require much less BCs to be inspected to pinpoint all faulty BCs in volatile environments. In large-scale scenarios, this advantage can lead to significant saving of system administrative cost. For example, for an SOS that consists of 100 services, our *c<sub>D</sub>-c<sub>0</sub>* approach requires an average of 3 to 13 less services to be inspected than *c<sub>0</sub>* ranking to localize all the anomalies. Our approaches can significantly speed up the anomaly detection process. Then, adaptation actions can be taken immediately to fix the system, relieving or eliminating the end-to-end system delay.

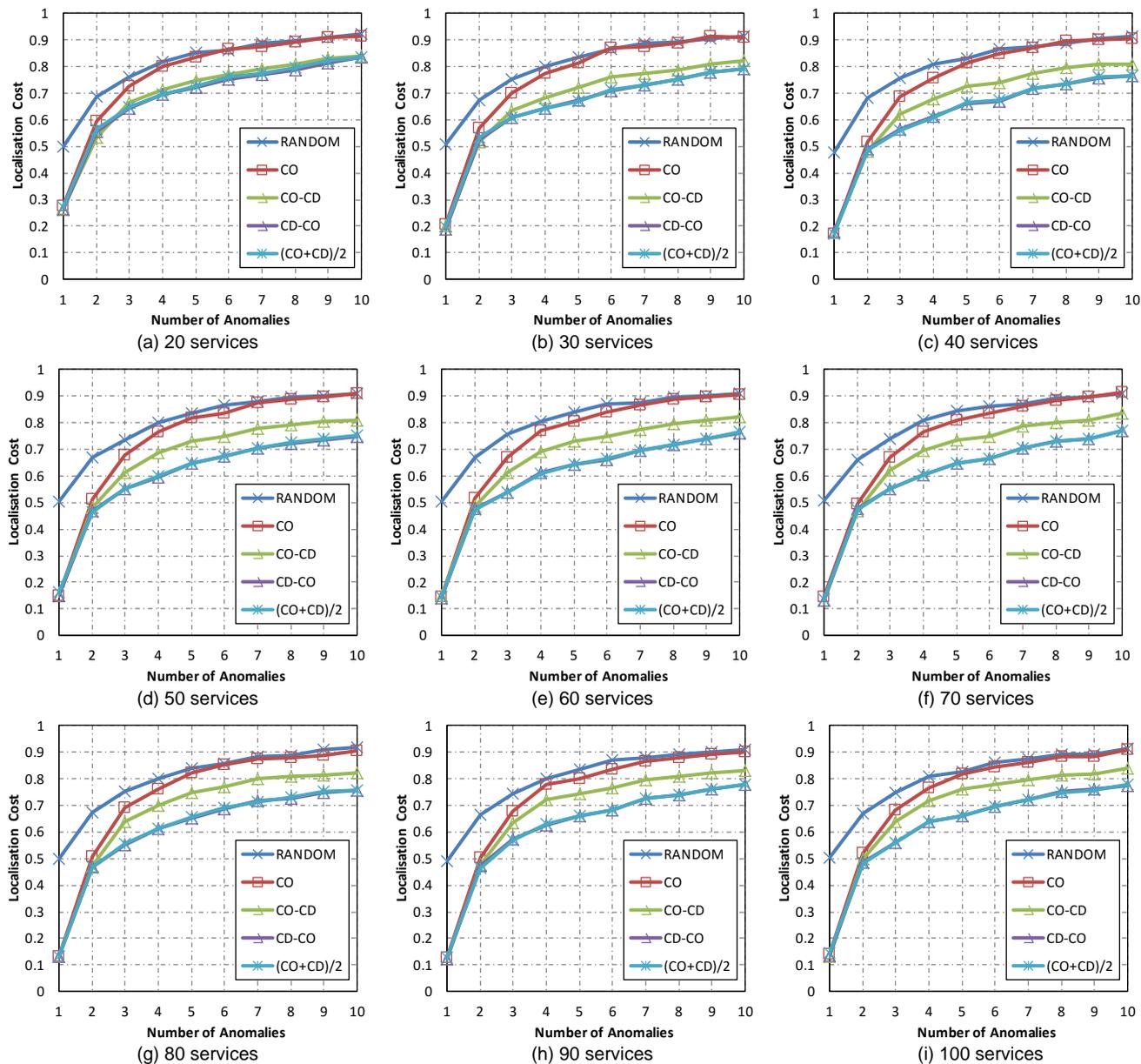


Fig. 6. Anomaly localization cost.

Although not explicitly demonstrated in Figure 5, it can be inferred that the localization costs of all approaches will reach 1.0 in an extreme environment where all BCs are faulty. In such a scenario, it does not make a difference which BCs to inspect first because all BCs are faulty.

Now we compare our own three approaches that rank BCs by their similarity coefficient scores and delay coefficient scores in different ways. The  $c_{D-CO}$  and  $(c_O+c_D)/2$  approaches demonstrate very similar performance, both outperforming the  $c_{O-CD}$  approach. The advantages of the  $c_{D-CO}$  and  $(c_O+c_D)/2$  approaches over the  $c_{O-CD}$  approach increase as the number of concurrent anomalies increase and as the size of the system increases. According to Table 4, the final winner is the  $c_{D-CO}$  approach, with only marginal advantage over the  $(c_O+c_D)/2$  approach.

To conclude, in a very stable operating environment, it does not make a big difference which anomaly localization approach is adopted. In a large-scale and volatile environment, our  $c_{D-CO}$  approach has the best perfor-

mance and is the most preferable choice over other approaches.

#### 4.2.2 Efficiency

As discussed in Section 1, the response time plays an important role in the success of an SOS. A runtime localisation approach must not take too much time to localize runtime anomalies occurring in an SOS because quick anomaly localization and rapid system adaptation are required to minimise the system delays perceived by the users of the system. An approach that can accurately rank the BCs according to their suspiciousness of being faulty is impractical if it needs to take a very long time. Here, we evaluate the efficiency of our approach, measured by the computation time taken to calculate the coefficients and rank the BCs, i.e., phases 4 and 5 in the anomaly localization procedure as presented in Fig. 3. The reason is that phases 1, 2 and 3 are performed offline at build-time while phases 4, 5 and 6 are performed at runtime and

directly contribute to the total time consumption for anomaly localization upon runtime anomalies.

Our experimental results show that even in the largest-scale scenarios with SOSs sized 100 services, the coefficient calculation and BC ranking take only a few milliseconds to complete. This is expected because the spectrum-based anomaly localization, which is a statistical technique, has long been acknowledged as an extremely lightweight approach [44]. Detailed experimental results are not presented in this subsection as it does not provide any new and insightful findings about spectrum-based anomaly localization. Compared to  $c_0$  ranking, our approach takes only less than 1 millisecond extra computation time. Such a short extra computation time is negligible in most, if not all, real-world scenarios.

Based on the above discussion, we conclude that our approach is highly efficient and can be used to localize runtime anomalies in all types of SOSs, including time-constrained SOSs, e.g., interactive and multimedia SOSs. We also analyse the threat to validity of our evaluation, as presented in Appendix A.

## 5 RELATED WORK

When operating in volatile environments, SOSs must be monitored in order to achieve timely and accurate detection and prediction of runtime anomalies. A great deal of monitoring techniques and approaches has been proposed for SOSs. Several languages are defined specifically for monitoring SOSs, such as WSCoL [7] and SECMOL [8]. These languages can be used to specify monitoring rules, constraints and assertions.

Many frameworks have also been proposed for monitoring SOSs. To name a few, Baresi et al. [9] propose a general and comprehensive solution for monitoring service compositions. In this monitoring solution, monitoring constraints can be defined on single and multiple instances, on punctual properties and on complete behaviours. Guinea et al. [15] propose a framework that integrates monitoring across the software and infrastructure layers. A variation of the MAPE control loop is introduced into the framework that acknowledges the multifaceted nature of SOSs. Kertész et al. [27] integrate SALMon [36] in IMA4SSP - their monitoring approach to seamless service provisioning - so that it can collect dynamic reliability information of SOSs expressed in pre-defined quality metrics.

Monitoring only the critical parts of an SOS is a cost-effective solution [17] to SOS monitoring. As a result, non-critical BCs will not be monitored. When an anomaly occurs in such a BC, it must be localized quickly so that adaptation actions can be taken to fix the anomaly immediately in order to avoid or reduce the delay caused to the system execution [7, 10, 19].

Many fault localization techniques have been proposed for different types of software systems. To name a few, Artzi et al. [6] present how fault localization algorithms can be enhanced to pinpoint faults effectively in Web applications. Park et al. [37] present a dynamic fault localization technique that can pinpoint faults in concurrent programs. Novotný et al. [39] present a method for localizing faults occurring in service-oriented systems on mo-

bile ad hoc networks. Sharma et al. [45] present a time-invariant relationships based approach for fault localization in distributed systems. Nguyen et al. [33] present FChain, a black-box online fault localization approach for diagnosing performance anomalies in cloud systems. It assumes that all the components of a cloud system are deployed within one cloud and thus all system metrics can be discovered immediately upon a detected performance anomaly. Thus, FChain is not suitable for SOSs whose component services are very often distributed across multiple clouds.

Compared to other fault localization techniques, including program slicing [49], delta debugging [31] and model-based diagnosis [29], spectrum-based fault localization is specifically designed to statistically calculate the likelihood of each software component being faulty based on information from passed and failed system runs [1]. It has been intensively studied in the context of traditional programming languages, such as C and Java, under the assumption that a component with a high similarity to the error vector has a higher probability of being the cause of the observed failure than a component with low similarity. There are many similarity coefficients that can be calculated to evaluate the suspiciousness of a software component being faulty [2]. The popular ones include the Jaccard coefficient, used by the Pinpoint tool [12], the coefficient used in the Tarantula fault localization tool [25] and the Ochiai coefficient  $c_0$  often used in the molecular biology domain [14]. Empirical performance comparison has been conducted among different similarity coefficients and the results showed that the Ochiai coefficient outperformed the Jaccard coefficient and the Tarantula coefficient [2, 44]. This has also been theoretically analysed and experimentally demonstrated by our previous work presented in [18] and [47]. Due to its high performance, Ochiai has been used in several faulty localization tools, e.g., Zoltar [22], GZoltar [11] and F<sup>3</sup> [23].

An SOS is different from a traditional standalone software system as it often operates in a distributed and volatile environment. Running test cases at build-time to localize the bugs in software systems, which is the method adopted by existing fault localisation techniques in the context of traditional programming languages, is impractical in localizing runtime anomalies occurring in SOSs. An SOS usually consists of multiple distributed component services, whose states are difficult to inspect upon the occurrence of runtime anomalies. However, the execution engine of the SOS, e.g., the BPEL engine, is centralised. The end-to-end response time of the SOS can be inspected easily and end-to-end system delays can be detected efficiently. Taking this advantage, this paper presents a fault localization approach to attack the challenging research problem of runtime anomaly localization for SOSs. Our approach evaluates BCs' suspiciousness of being faulty by calculating their similarity coefficient scores and delay coefficient scores. The BCs that are more likely to be faulty are prioritised in the inspection.

## 6 CONCLUSIONS

In this paper, we propose a spectrum-based approach for anomaly localization in service-oriented systems. The

main idea is to exploit end-to-end system quality data for localizing runtime anomalies. The system spectra are analysed to calculate the basic components' suspiciousness of resulting in the runtime anomaly. BCs' similarity coefficient scores and delay coefficient scores are calculated, which are used to rank the BCs. The comprehensive experimental analysis shows the effectiveness and efficiency of our approach, and proves that our approach can significantly save the time and efforts for runtime anomaly localization.

As part of future work, we plan to experiment with additional similarity metrics that can be used to rank BCs, and evaluate how their effectiveness compares to the approach presented in this paper. In addition, we will also investigate the statistical relation between localization cost and the number of execution scenarios.

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